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# Notes on the estimation of the asymptotics of the moments for the m collector's problem

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#### Abstract

The general collector's problem describes a process in which N distinct coupons are placed in an urn and randomly selected one at a time (with replacement) until at least m of all N existing different types of coupons have been selected. Let  $T_m(N)$  the random variable denoting the number of trials needed for this goal. We briefly present the leading asymptotics of the (rising) moments of  $T_m(N)$  as  $N \to \infty$  for large classes of coupon probabilities. It is proved that the expectation of  $T_m(N)$  becomes minimum when the coupons are uniformly distributed. Moreover, a theorem on the asymptotic estimates of the rising moments of  $T_m(N)$  by comparison with known sequences of coupon probabilities is proved.

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#### **1** Introduction

We consider the following classical urn problem. Suppose that N distinct types of balls are placed in an urn from which balls are being collected independently with replacement, each one with probability  $p_j$ ,  $j = 1, 2, \dots, N$ . Let  $T_m(N)$  be the number of trials needed until each ball has been collected mtimes, where m is a fixed positive integer. This process is, sometimes, called double dixie cup problem, while for the particular case where m = 1 is the socalled coupon collector's problem. The problem for the case m = 1 has a long history (its origin can be traced back to De Moivre's treatise De Mensura Sortis of 1712 and Laplace's pioneering work Theorie Analytique de Probabilites of 1812), and its applications lie on several areas of science hence (e.g., biology, linguistics, search algorithms). For general values of m and for  $p_j = 1/N$  D. J. Newman and L. Shepp [8] and soonafter, P. Erdős and A. Rényi [6] determined the expectation, as well as the limit distribution of  $T_m(N)$ . They proved that

$$\lim_{N \to \infty} P\left\{\frac{T_m(N) - N\ln N - (m-1)N\ln\ln N + N\ln(m-1)!}{N} \le y\right\} = e^{-e^{-y}}.$$
(1.1)

For general values of m and for the case of unequal coupon probabilities one nay find useful results in [4], where the authors developed techniques of computing the asymptotics of the first and the second moment of  $T_m(N)$ , the variance, as well as, the limit distribution for large classes of coupon probabilities.Let

$$T_m(N)^{(r)} := T_m(N)(T_m(N)+1)\cdots(T_m(N)+r-1), \qquad r = 1, 2, \dots$$
 (1.2)

i.e., r-th rising moment of  $T_m(N)$ . In this paper we present leading asymptotics for the rising moments of the random variable  $T_m(N)$ , for rich classes of probabilities. We prove that  $E[T_m(N)^{(r)}]$  becomes minimum when the  $p_j$ 's are uniformly distributed by using the Schur - Ostrowski criterion. Finally, a theorem that helps us obtain asymptotic estimates by comparison with sequences of coupon probabilities, for which the asymptotics are known, is presented.

### **2** The rising moments of $T_m(N)$

Let  $\alpha = \{a_j\}_{j=1}^{\infty}$  be a sequence of strictly positive numbers. Then, for each integer N > 0, one can create a probability measure  $\pi_N = \{p_1, ..., p_N\}$  on the

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set of types  $\{1,...,N\}$  by taking

$$p_j = \frac{a_j}{A_N}, \quad \text{where} \quad A_N = \sum_{j=1}^N a_j.$$
 (2.1)

By a Poissonization technique it is not hard to get explicit formulae for the moments and the moment generating function of  $T_m(N)$  (see, [4]):

$$E\left[T_m(N)^{(r)}\right] = r \int_0^\infty \left\{ 1 - \prod_{j=1}^N \left[1 - S_m(p_j t)e^{-p_j t}\right] \right\} t^{r-1} dt \qquad (2.2)$$

$$G(z) := E\left[z^{-T_m(N)}\right]$$

$$= 1 - (z-1) \int_0^\infty \left\{ 1 - \prod_{j=1}^N \left[1 - S_m(p_j t) e^{-p_j t}\right] \right\} e^{-(z-1)t} dt,$$
(2.3)

for  $\Re(z) > 1$ , r = 1, 2, ..., and  $S_m(y)$  denotes the *m*-th partial sum of  $e^y$ , namely

$$S_m(y) := 1 + t + \frac{y^2}{2!} + \dots + \frac{y^{m-1}}{(m-1)!} = \sum_{l=0}^{m-1} \frac{y^l}{l!} \,. \tag{2.4}$$

We introduce the notation

$$E_m(N;\alpha;r) := r \int_0^\infty \left[ 1 - \prod_{j=1}^N \left( 1 - e^{-a_j t} S_m(a_j t) \right) \right] t^{r-1} dt.$$
 (2.5)

For a sequence  $\alpha = \{a_j\}_{j=1}^{\infty}$  and a number s > 0 we set  $s\alpha = \{sa_j\}_{j=1}^{\infty}$ . Hence,

$$E\left[T_m(N)^{(r)}\right] = A_N^r E_m(N;\alpha;r).$$
(2.6)

Under (2.6) the problem of estimating  $E\left[T_m(N)^{(r)}\right]$  can be treated as two separate problems, namely estimating  $A_N^r$  and estimating  $E_m(N;\alpha; r)$ , (see (2.5)). The estimation of  $A_N^r$  can be considered an external matter which can be handled by existing powerful methods, such as the Euler-Maclaurin sum formula, the Laplace method for sums (see, e.g., [1]), or even summation by parts. Let

$$L_{m}(N;\alpha;r) := \lim_{N} E_{m}(N;\alpha;r)$$

$$= r \int_{0}^{\infty} \left[ 1 - \prod_{j=1}^{\infty} \left( 1 - e^{-a_{j}t} S_{m}(a_{j}t) \right) \right] t^{r-1} dt.$$
(2.7)

**Theorem 2.1.** For any fixed positive integers m and r,  $E\left[T_m(N)^{(r)}\right]$  becomes minimum when all  $p_j$ 's are equal.

*Proof.* To prove the theorem it suffices to show that, for a fixed t > 0, the maximum of the quantity

$$\prod_{j=1}^{N} \left[ 1 - e^{-p_j t} S_m(p_j t) \right],$$

subject to the constraints  $p_1 + \cdots + p_N = 1$ ,  $p_j > 0$ ,  $j = 1, 2, \cdots, N$ , occurs when all  $p_j$ 's are equal. Set  $\phi : (0, 1)^N \longrightarrow (0, \infty)$ ,

$$\phi(p_1 \cdots, p_N) := \sum_{j=1}^N \ln\left[1 - e^{-p_j t} S_m(p_j t)\right].$$
(2.8)

Clearly,  $\phi$  is symmetric w.r.t. its variables. Now, if for all  $1 \le i \ne j \le N$ ,

$$(p_i - p_j) \left( \frac{\partial \phi (p_1, p_2, \cdots, p_N)}{\partial p_i} - \frac{\partial \phi (p_1, p_2, \cdots, p_N)}{\partial p_j} \right) \le 0, \qquad (2.9)$$

then,  $\phi$  will be a Schur–concave function (see, [7], page 84, theorem A.4) and will attain its maximum when all  $p_j$ 's are equal (see, [7], page 413). We have

$$\frac{\partial \phi\left(p_{1}, p_{2}, \cdots, p_{N}\right)}{\partial p_{i}} = \frac{t}{(m-1)!} \cdot \frac{e^{-p_{i}t}\left(tp_{i}\right)^{m-1}}{1 - e^{-p_{i}t}S_{m}\left(p_{i}t\right)}$$

It suffices to obtain that the function  $f(\cdot)$  is decreasing, where

$$f(x) := \frac{e^{-x}x^{m-1}}{1 - e^{-x}S_m(x)}, \ x > 0.$$

Observing that

$$(e^{-x}S_m(x))' = -\frac{e^{-x}x^{m-1}}{(m-1)!},$$

we have

$$f'(x) = \frac{e^{-x}x^{m-2}}{\left[1 - e^{-x}S_m(x)\right]^2} g(x), \qquad (2.10)$$

where

$$g(x) := (m - 1 - x) \left[ 1 - e^{-x} S_m(x) \right] - \frac{e^{-x} x^m}{(m - 1)!}.$$
 (2.11)

Notice that g(x) extends to a smooth function on  $\mathbb{R}$ . In particular g(0) = 0. If m = 1, then g(x) = -x and (2.11) implies that f'(x) < 0 for all x > 0. For  $m \ge 2$  we have

$$g'(x) = -1 + e^{-x}S_m(x) - \frac{e^{-x}x^{m-1}}{(m-1)!}, \ g'(0) = 0,$$

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and

$$g''(x) = -(m-1)\frac{e^{-x}x^{m-2}}{(m-1)!} < 0, \qquad x > 0.$$

Thus g(x) < 0 for all x > 0. Therefore, f'(x) < 0 for all x > 0 and the proof is completed.

The following theorem is related to our recent work [4] and the proof is omitted.

**Theorem 2.2.**  $L_m(N; \alpha; r) < \infty$  simultaneously for all positive (fixed) integers m and r, if and only if there exist a  $\xi \in (0, 1)$  such that

$$\sum_{j=1}^{\infty} \xi^{a_j} < \infty$$

If  $L_m(N; \alpha; r) < \infty$ , then for all positive integers m and r we have

$$E\left[T_m(N)^{(r)}\right] = A_N^r L_m(N;\alpha;r) \left[1 + o(1)\right] \quad as \ N \to \infty.$$

Examples of this case are the positive power law, namely  $\alpha = \{j^p\}_{j=1}^{\infty}$ , where p > 0. In particular, when p = 1 we have the so-called *linear* case. Also, the families of sequences  $\kappa = \{e^{qj}\}_{j=1}^{\infty}$  and where q > 0 fall in this case. Notice that the sequences  $\beta = \{e^{-qj}\}_{j=1}^{\infty}$  produce the same coupon probabilities with  $\kappa$ , hence they are covered too.

For the challenging case where  $L_m(N; \alpha; r) = \infty$  for some fixed positive integer r (and for any fixed m) we write  $a_i$  in the form

$$a_j = f(j)^{-1} (2.12)$$

where

$$f(x) > 0$$
 and  $f'(x) > 0$ , (2.13)

and we will discuss our problem for large classes of distributions. In particular, we will cover the cases where  $f(\cdot)$  belongs to the class of positive and strictly increasing  $C^3(0,\infty)$  functions, which grow to  $\infty$  (as  $x \to \infty$ ) slower than exponentials, but faster than powers of logarithms. We assume that f(x) possesses three derivatives satisfying the following conditions as  $x \to \infty$ :

(i) 
$$f(x) \to \infty$$
,  
(ii)  $\frac{f'(x)}{f(x)} \to 0$ ,  
(iii)  $\frac{f''(x)/f'(x)}{f'(x)/f(x)} = O(1)$ ,  
(iv)  $\frac{f'''(x) f(x)^2}{f'(x)^3} = O(1)$ .  
(2.14)

These conditions are satisfied by a variety of commonly used functions. For example,

$$f(x) = x^p (\ln x)^q, \quad p > 0, \ q \in \mathbb{R}, \qquad f(x) = \exp(x^r), \quad 0 < r < 1,$$

or various convex combinations of products of such functions. An important example falling in this case is the well known generalized Zipf law, namely  $f(x) = x^p$ , where p > 0. Zipf's law has attracted the interest of scientists of several areas of science, such as linguistics, biology, etc.

With similar arguments as in [4] one has the following theorem for the rising moments of the random variable  $T_m(N)$ .

**Theorem 2.3.** If  $\alpha = \{1/f(j)\}_{j=1}^{\infty}$ , where  $f(\cdot)$  satisfies (2.13) and (2.14), then as  $N \to \infty$ 

$$E\left[T_N^{(r)}\right] \sim \frac{1}{\min_{1 \le j \le N} \{p_j\}^r} \ln\left(\frac{f(N)}{f'(N)}\right)^r.$$
(2.15)

## 3 Asymptotic estimates for the rising moments of $T_N$ by comparison with known sequences

Here we will present a theorem that helps us obtain asymptotic estimates by comparison with sequences  $\alpha$  for which the asymptotic estimates of  $E_m(N; \alpha; r)$ are known (for instance, via Theorem 2.3). First, we recall the following notation. Suppose that  $\{s_j\}_{j=1}^{\infty}$  and  $\{t_j\}_{j=1}^{\infty}$  are two sequences of nonnegative terms. The symbol  $s_j \simeq t_j$  means that there are two constants  $C_1 > C_2 > 0$ and an integer  $j_0 > 0$  such that

$$C_2 t_j \le s_j \le C_1 t_j, \qquad \text{for all } j \ge j_0, \tag{3.1}$$

i.e.  $s_j = O(t_j)$  and  $t_j = O(s_j)$ .

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**Theorem 3.1.** Let  $\alpha = \{a_j\}_{j=1}^{\infty}$  and  $\beta = \{b_j\}_{j=1}^{\infty}$  be sequences of strictly positive terms such that  $\lim_N E_m(N; \alpha; r) = \lim_N E_m(N; \beta; r) = \infty$ . (i) If there exists an  $j_0$  such that  $a_j = b_j$ , for all  $j \ge j_0$ , then  $E_m(N; \alpha; r) - E_m(N; \beta; r)$  is bounded, (ii) if  $a_j = O(b_j)$ , then  $E_m(N; \beta; r) = O(E_m(N; \alpha; r) \text{ as } N \to \infty,$ (iii) if  $a_j = o(b_j)$ , then  $E_m(N; \beta; r) = o(E_m(N; \alpha; r))$  as  $N \to \infty$ , (iv) if  $a_j \approx b_j$ , then  $E_m(N; \beta; r) \approx E_m(N; \alpha; r)$  as  $N \to \infty$ , (v) if  $a_j \sim b_j$ , then  $E_m(N; \beta; r) \sim E_m(N; \alpha; r)$  as  $N \to \infty$ .

*Proof.* Case (i) follows easily from (2.5):

 $|E_m(N;\alpha;r) - E_m(N;\beta;r)| =$ 

$$= r \left| \int_{0}^{\infty} \prod_{j=j_{0}}^{N} \left( 1 - S_{m}(a_{j}t)e^{-a_{j}t} \right) \right|^{j_{0}-1} \left( 1 - S_{m}(b_{j}t)e^{-b_{j}t} \right) \left| t^{r-1} dt \right|$$

$$\leq r \int_{0}^{\infty} \left| \left[ \prod_{j=1}^{j_{0}-1} \left( 1 - S_{m}(a_{j}t)e^{-a_{j}t} \right) - \prod_{j=1}^{j_{0}-1} \left( 1 - S_{m}(b_{j}t)e^{-b_{j}t} \right) \right] \right| t^{r-1} dt$$

$$= r \int_{0}^{\infty} \left| \sum_{J \subset \{1, \dots, j_{0}-1\}} \left( -1 \right)^{|J|} \left\{ \exp \left( -t \sum_{j \in J} a_{j} \right) \prod_{j \in J} S_{m}(a_{j}t) - \exp \left( -t \sum_{j \in J} b_{j} \right) \prod_{j \in J} S_{m}(b_{j}t) \right\} t^{r-1} dt < \infty,$$

where we have used the formula

$$\prod_{j=1}^{N} \left( 1 - S_m(p_j t) e^{-p_j t} \right) = \sum_{J \subset \{1,\dots,N\}} \left( -1 \right)^{|J|} \exp\left( -t \sum_{j \in J} p_j \right) \prod_{j \in J} S_m(p_j t). \quad (3.2)$$

Notice that the sum extends over all  $2^{j-1}$  subsets J of  $\{1, ..., j-1\}$ , while |J| denotes the cardinality of J.

(ii) Since  $a_j = O(b_j)$ , there is a positive constant M and an integer  $j_0$ , such that  $a_j \leq Mb_j$ , for all  $j \geq j_0$ . By part (i) of the theorem we have

$$|E_m(N; M\beta; r) - E_m(N; \alpha; r)| \le C,$$

for some positive constant C as  $N \to \infty$ . Next observe that (2.5) implies

$$E_m(N; s\alpha; r) = s^{-r} E_m(N; \alpha; r).$$
(3.3)

Using (3.3) we get

$$\left|\frac{1}{M^r}E_m(N;\beta;r) - E_m(N;\alpha;r)\right| \le C,$$

i.e.

$$E_m(N;\beta;r) \le M^r E_m(N;\alpha;r) + CM^r,$$

and the result follows immediately from the definition of the O notation.

(iii) Fix an  $\epsilon > 0$ . Then  $a_j \leq \epsilon b_j$ , for all  $j \geq j_0(\epsilon)$ . Thus, by part (i) there is an  $M = M(\epsilon)$  such that

$$E_m(N;\epsilon\beta;r) - E_m(N;\alpha;r) \le M.$$

By invoking (3.3) we get

$$\frac{1}{\epsilon^r} E_m(N;\beta;r) \le E_m(N;\alpha;r) + M, \quad \text{for all } N \ge N_0(\epsilon).$$

If we divide by  $E_m(N; \alpha; r)$  and then let  $N \to \infty$ , we obtain (iii), since  $\epsilon$  is arbitrary and  $\lim_N E_m(N; \alpha; r) = \infty$ .

(iv) Since  $a_j \simeq b_j$ , then from (3.1) we have  $a_j = O(b_j)$  and  $b_j = O(a_j)$ . Using part (ii) we get as  $N \to \infty$ ,  $E_m(N;\beta;r) = O(E_m(N;\alpha;r))$  and  $E_m(N;\alpha;r) = O(E_m(N;\beta;r))$ , the result follows again from (3.1).

To prove (v) we first fix an  $\epsilon > 0$ . Then  $(1 - \epsilon)b_j \leq a_j \leq (1 + \epsilon)b_j$ , for all  $j \geq j_0(\epsilon)$ . Thus, by case (i) and (3.3) there is an  $M = M(\epsilon)$  such that

$$\left(\frac{1}{1+\epsilon}\right)^r E_m(N;\beta;r) - M \le E_m(N;\alpha;r) \le \left(\frac{1}{1-\epsilon}\right)^r E_m(N;\beta;r) + M,$$

for all  $N \ge N_0(\epsilon)$ . If we divide by  $E_m(N;\beta;r)$  and then let  $N \to \infty$ , we obtain (v) since  $\epsilon$  is arbitrary and  $\lim_N E_m(N;\beta;r) = \infty$ .

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